

Chapter 3: TRENDS.  $\rightarrow$  Deterministic  
(Regression).

7/11/12 (1)

Chapter: Models for stationary  
Time Series.

General Linear Process

$$Y_t = e_t + \psi_1 e_{t-1} + \psi_2 e_{t-2} + \dots$$

We also assume

$$\sum_{k=1}^{\infty} \psi_k^2 < \infty$$

can assume  $\psi_0 = 1$ .

Example: Suppose  $\psi_k = \phi^k$

$$Y_t = e_t + \phi e_{t-1} + \phi^2 e_{t-2} + \phi^3 e_{t-3} + \dots$$

$$\begin{aligned} \mathbb{E}Y_t &= \mathbb{E}(e_t + \phi e_{t-1} + \phi^2 e_{t-2} + \dots) \\ &= 0 + \phi \cdot 0 + \phi^2 \cdot 0 + \dots \\ &= 0. \end{aligned}$$

Variance:  $\gamma_0$

$$\begin{aligned} &\text{Var}(e_t + \phi e_{t-1} + \phi^2 e_{t-2} + \dots) \\ &= \text{Var}(e_t) + \text{Var}(\phi e_{t-1}) + \text{Var}(\phi^2 e_{t-2}) + \dots \\ &= \sigma^2 + \phi^2 \cdot \sigma^2 + \phi^4 \cdot \sigma^2 + \dots \\ &= \sigma^2 (1 + \phi^2 + \phi^4 + \dots) \\ &= \sigma^2 \cdot \sum_{k=0}^{\infty} [\phi^2]^k \end{aligned}$$

$$= \sigma^2 \cdot \frac{1}{1 - \phi^2}, \quad |\phi^2| < 1$$

$$-1 < \phi < 1.$$

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$$\gamma_1 = \text{Cov}(Y_t, Y_{t-1})$$

$$= \text{Cov}(e_t + \phi e_{t-1} + \phi^2 e_{t-2} + \dots, e_{t-1} + \phi e_{t-2} + \phi^2 e_{t-3} + \dots)$$

$$= 0 + \phi \sigma^2 + \phi^3 \sigma^2 + \phi^5 \sigma^2 + \phi^7 \sigma^2 \dots$$

$$= \phi \sigma^2 (1 + \phi^2 + \phi^4 + \dots)$$

$$= \frac{\phi \sigma^2}{1 - \phi^2}$$

Then autocorrelation:

$$\rho_1 = \frac{\gamma_1}{\gamma_0} = \frac{\cancel{\phi \sigma^2} / (1 - \phi^2)}{\cancel{\sigma^2} / (1 - \phi^2)}$$

$$= \phi$$

Similarly,

$$\rho_h = \phi^h$$

Does NOT depend on  $t$ !  
STATIONARY!

For general linear:

$$\gamma_h = \sigma_e^2 \sum_{i=0}^{\infty} \psi_i \psi_{i+h}$$

Moving Average Process:

only finitely many of  $\psi$ 's non-zero.

Example: MA(1)

$$Y_t = e_t + \theta e_{t-1}, \quad t = 1, 2, 3, \dots$$

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Then

$$\begin{aligned}\mathbb{E} I_t &= \mathbb{E}(e_t + \theta e_{t-1}) \\ &= 0 + 0 = 0.\end{aligned}$$

$$\begin{aligned}\text{Var}(I_t) &= \text{Var}(e_t + \theta e_{t-1}) \\ &= \sigma^2 + \theta^2 \sigma^2 \\ &= \sigma^2(1 + \theta^2) \\ &= \gamma_0\end{aligned}$$

$$\gamma_1 = \theta \cdot \sigma^2$$

which means

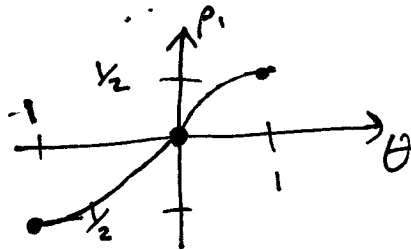
$$\begin{aligned}\rho_1 &= \frac{\gamma_1}{\gamma_0} \\ &= \frac{\theta \sigma^2}{\sigma^2(1 + \theta^2)}\end{aligned}$$

$$\rho_1 = \frac{\theta}{1 + \theta^2}$$

and

$$\rho_{k2} = 0 \text{ otherwise...}$$

$$\rho = \theta / (1 + \theta^2)$$

How does  $\rho$  change with  $\theta$ ?WATCH THIS:What if plug in  $\frac{1}{\theta}$ ?

$$\rho = \frac{\theta}{1 + \theta^2}$$

$$\begin{aligned}&\frac{(\frac{1}{\theta})}{1 + (\frac{1}{\theta})^2} \\ &= \frac{\frac{1}{\theta}}{1 + \frac{1}{\theta^2}} \cdot \frac{\theta^2}{\theta^2}\end{aligned}$$

$$= \frac{\theta}{\theta^2 + 1}$$

That is,

MA(1) with  $\theta = 0.7$

has same autocorrelation as

MA(1) with  $\theta = 10/7$

Invertibility

Ex] Plots.

Next: MA(2)

$$\underline{Y}_t = e_t + \theta_1 e_{t-1} + \theta_2 e_{t-2}$$

$$\begin{aligned} \text{Mean} &= E(e_t + \theta_1 e_{t-1} + \theta_2 e_{t-2}) \\ &= 0 + 0 + 0 \\ &= 0 \end{aligned}$$

$$\begin{aligned} \text{Var}(\underline{Y}_t) &= V(e_t + \theta_1 e_{t-1} + \theta_2 e_{t-2}) \\ &= \sigma^2 + \theta_1^2 \sigma^2 + \theta_2^2 \sigma^2 \\ &= \sigma^2 (1 + \theta_1^2 + \theta_2^2) \end{aligned}$$

$$\gamma_1 = \theta_1 + \theta_1 \theta_2 \sigma^2$$

$$\gamma_2 = \theta_2 \sigma^2$$

$$\gamma_3 = 0 \dots$$

And so autocorrelation

$$\rho_1 = \frac{\gamma_1}{\gamma_0}, \quad \rho_2 = \frac{\gamma_2}{\gamma_0}, \quad \rho_k = 0 \text{ otherwise}$$

## General MA(q) process

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$$Y_t = e_t + \theta_1 e_{t-1} + \theta_2 e_{t-2} + \dots + \theta_q e_{t-q}$$

$$\text{Var}(Y_t) = \sigma^2 (1 + \theta_1^2 + \theta_2^2 + \dots + \theta_q^2)$$

Auto correlation:

$$\rho_h = \frac{\theta_h + \theta_1 \theta_{h+1} + \dots + \theta_{q-h} \theta_q}{1 + \theta_1^2 + \dots + \theta_q^2}, \quad h=1, \dots, q$$

and  $\rho_q$  has numerator  $\theta_q$ .

and  $\rho_h = 0$  otherwise.  
"Cuts off"

Autoregressive Process: regression on itself.

- depends on past values of series plus "innovation"

Ex] AR(1)

$$Y_t = \phi Y_{t-1} + e_t$$

$$\text{Mean: } E(Y_t) = E(\phi Y_{t-1} + e_t)$$

$$\mu = \phi \mu + 0$$

$$\text{Variance: } \text{Var}(Y_t) = \text{Var}(\phi Y_{t-1} + e_t)$$

$$\sigma_y^2 = \phi^2 \sigma_y^2 + \sigma_e^2$$

$$\sigma_y^2 - \phi^2 \sigma_y^2 = \sigma_e^2$$

$$\sigma_y^2 (1 - \phi^2) = \sigma_e^2$$

$$\sigma_y^2 = \frac{\sigma_e^2}{1 - \phi^2}$$

Autocovariance:

$$Y_{t-k} - \bar{Y} = \phi(Y_{t-k-1} - \bar{Y}) + e_{t-k}$$

$$E(\dots) = E(\dots) + E(\dots)$$

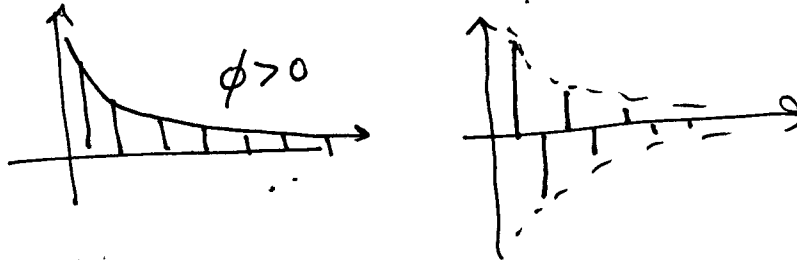
$$\gamma_k = \phi \gamma_{k-1} + 0$$

$$\text{Thus } \gamma_k = \phi^k \cdot \frac{\sigma^2}{1-\phi^2}, \quad k=1, \dots$$

$$\text{And } \rho_k = \phi^k, \quad k=1, 2, 3, \dots$$

STATIONARY.

Make plots of ACF



Stationarity condition: need  
 $|\phi| < 1$ .

General Linear process:

$$\begin{aligned} Y_t &= \phi Y_{t-1} + e_t \\ &= \phi(\phi Y_{t-2} + e_{t-1}) + e_t \\ &= \phi(\phi(\phi Y_{t-3} + e_{t-2}) + e_{t-1}) + e_t \\ &\vdots \\ &= e_t + \phi e_{t-1} + \phi^2 e_{t-2} + \dots \end{aligned}$$

Next: AR(2)

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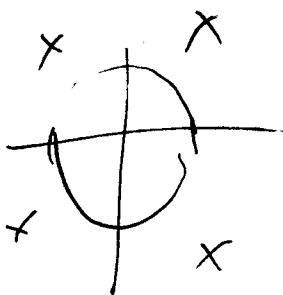
$$\Psi_t = \phi_1 \Psi_{t-1} + \phi_2 \Psi_{t-2} + \epsilon_t$$

"Characteristic Polynomial"

$$\phi(x) = 1 - \phi_1 x - \phi_2 x^2$$

"Characteristic Equation"

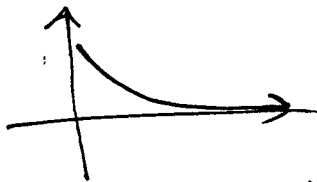
$$\phi(x) = 0$$



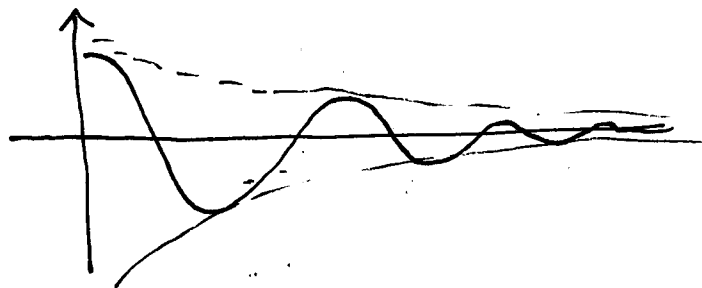
- Stationarity condition:  
Need roots of CE to be  
greater than 1 in modulus.

- ACF: complicated. Depends  
on roots of CE.

Shape: If real roots,  
ACF goes to zero exponentially



If roots <sup>of CE</sup> are complex, looks  
like "damped sine function"



## Yule-Walkers Equations:

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$$Y_{t_2} = \phi_1 Y_{t_2-1} + \phi_2 Y_{t_2-2}$$

$$t_2 = 2, 3, 4, \dots$$

General Linear process

MORE COMPLICATED,  
depends on roots of CE.

General AR(p)

$$Y_t = \phi_1 Y_{t-1} + \phi_2 Y_{t-2} + \dots + \phi_p Y_{t-p} + e_t$$

Characteristic poly:

$$\phi(x) = 1 - \phi_1 x - \phi_2 x^2 - \dots - \phi_p x^p$$

Char. Eq<sup>n</sup>.

$$\phi(x) = 0$$

All roots must be greater than 1  
in modulus to be stationary AR(p)

A.K.A. "causal model".

Necessary conditions.

$$\sum_{k=1}^p \phi_k < 1 \quad \text{and} \quad |\phi_p| < 1$$

Mixed Auto regressive Moving Average  
Models

ARMA(p, q)

$$Y_t = \phi_1 Y_{t-1} + \phi_2 Y_{t-2} + \dots + \phi_p Y_{t-p} + e_t + \dots \\ \dots + \theta_1 e_{t-1} + \theta_2 e_{t-2} + \dots + \theta_q e_{t-q}$$



Stationarity: need roots of AR polynomial fall outside unit circle.

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Also need: no common factors in AR  
M.A. characteristic polynomial.

AR model:  $\phi(x) = 1 - \phi_1 x - \dots - \phi_p x^p$ .

MA charact. poly nomial

$$\theta(x) = 1 + \theta_1 x + \theta_2 x^2 + \dots + \theta_q x^q$$

Example:

$$Y_t = e_t$$

White  
noise \*

$$Y_{t-1} = e_{t-1}$$

$$0.7 Y_{t-1} = 0.7 e_{t-1} \quad **$$

Then what is

\* - \*\*

$$Y_t - 0.7 Y_{t-1} = e_t - 0.7 e_{t-1}$$

$$Y_t = 0.7 Y_{t-1} + e_t - 0.7 e_{t-1}$$

this looks like

ARMA(1, 1)!!

But look at char. polynomials:

AR:  $\phi(x) = 1 - 0.7x$

MA:  $\theta(x) = 1 - 0.7x$

SAME!

Example: ARMA(1,1)

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$$Y_t = \phi Y_{t-1} + e_t + \theta e_{t-1}$$

Always assume no common factors for polyn. which means  $\phi \neq \theta$

Variance:

$$\gamma_0 = \sigma_e^2 (1 + 2\phi\theta + \theta^2) / (1 - \phi^2)$$

Auto correlation:

$$\rho_h = \frac{\phi^{h-1} (1 + \theta\phi)(\phi + \theta)}{1 + 2\theta\phi + \theta^2}$$

General linear process

$$Y_t = e_t + (\phi + \theta) \sum_{j=1}^{\infty} \phi^{j-1} e_{t-j}$$

Fact: MA models can be nonunique in general.

KNOW: Can rewrite AR process as linear process.

Ex) KNOW AR(1):

$$Y_t = e_t + \underbrace{\phi}_{\theta_1} e_{t-1} + \underbrace{\phi^2}_{\theta_2} e_{t-2} + \dots$$

Conclusion: AR(p) is just an MA( $\infty$ ) process!

Question: Can you start with

$MA(q)$ , jumble around, get

$AR(\infty)$  ????

Yes!!! (sometimes).

Look at  $MA(q)$  characteristic polynomial

$$\theta(x) = 1 + \theta_1 x + \theta_2 x^2 + \dots + \theta_q x^q$$

We say an  $MA(q)$  model is

"INVERTIBLE" iff

all roots of  $\theta(x) = 0$  to  
fall ~~outside~~ unit circle.

Invertible models are  
UNIQUE for given  
autocorrelation

Exercises:

Undergrad 2-5, 7-10

Grad: 12, 15, 18, 23

Note: check out

ARMAacf function in TSA  
package.

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